

Solution 3:

2. 由于独立 $\bar{X} - \bar{Y} \sim N(0, \frac{2\sigma^2}{n})$ 其分布函数记为 $\Phi_{0, \frac{2\sigma^2}{n}}(x)$

$$P(|\bar{X} - \bar{Y}| > 6) \approx 0.01 \Leftrightarrow P(|\bar{X} - \bar{Y}| \leq 6) \approx 0.99$$

$$P(|\bar{X} - \bar{Y}| \leq 6) = 2\Phi_{0, \frac{2\sigma^2}{n}}(6) - 1 = 2\Phi\left(\frac{6}{\sqrt{\frac{2\sigma^2}{n}}}\right) - 1 \approx 0.99$$

$$\Phi\left(\frac{6}{\sqrt{\frac{2\sigma^2}{n}}}\right) \approx 0.995 \Rightarrow \sqrt{\frac{n}{2}} \approx 2.58 \Rightarrow n \approx 13.32$$

$\therefore n$ 取 14.

□

11. 由于独立, $X_{n+1} - \bar{X} \sim N(0, \frac{n+1}{n}\sigma^2)$

则有
$$\frac{X_{n+1} - \bar{X}}{\sqrt{\frac{n+1}{n}\sigma^2}} \sim N(0, 1)$$

注意到
$$S_n^2 = \frac{n S_n^2}{n-1} \quad (S^2 = \frac{1}{n-1} \sum (x_i - \bar{x})^2)$$

由 $\frac{(n-1)S_n^2}{\sigma^2} \sim \chi_{n-1}^2$ 得 $\frac{n S_n^2}{\sigma^2} \sim \chi_{n-1}^2$

由分布定义, 及 $\bar{X}$ 与 S_n^2 独立, 而 $\bar{X}$ 与 S_n^2 独立可得

$\frac{X_{n+1} - \bar{X}}{\sqrt{\frac{n+1}{n}\sigma^2}}$ 与 $\frac{n S_n^2}{\sigma^2}$ 独立

$$\Rightarrow \frac{\frac{X_{n+1} - \bar{X}}{\sqrt{\frac{n+1}{n}\sigma^2}}}{\sqrt{\frac{n S_n^2}{\sigma^2}} / \sqrt{n-1}} = \frac{X_{n+1} - \bar{X}}{S_n} \sqrt{\frac{n-1}{n+1}} \sim t_{n-1}$$

□

12. $\bar{X} - \mu_1 \sim N(0, \frac{\sigma^2}{m})$ $\bar{Y} - \mu_2 \sim N(0, \frac{\sigma^2}{n})$ 由 $\bar{X} - \mu_1$ 与 $\bar{Y} - \mu_2$ 独立得

$$Z = \alpha(\bar{X} - \mu_1) + \beta(\bar{Y} - \mu_2) \sim N(0, (\frac{\alpha^2}{m} + \frac{\beta^2}{n})\sigma^2)$$

$$\Rightarrow \frac{Z}{\sqrt{(\frac{\alpha^2}{m} + \frac{\beta^2}{n})\sigma^2}} \sim N(0, 1)$$

又由 $S_{X+Y}^2 = \frac{m S_{1m}^2 + n S_{2n}^2}{m+n-2} \sim \chi_{m+n-2}^2$ 且 $\bar{X}$ 与 S_{1m}^2 独立, $\bar{Y}$ 与 S_{2n}^2 独立, S_{1m}^2 与 S_{2n}^2 独立

$$\Rightarrow \frac{Z}{\sqrt{(\frac{\alpha^2}{m} + \frac{\beta^2}{n})\sigma^2}} \text{ 与 } \frac{S_{X+Y}^2}{\sigma^2} \text{ 独立.}$$

$\therefore$ 由分布定义, $T \sim t_{m+n-2}$.

□



13. $X_1, \dots, X_n \text{ iid } \sim N(a, \sigma^2)$

$$\text{令 } \begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{pmatrix} = \begin{pmatrix} \frac{1}{\sqrt{n}} & \frac{1}{\sqrt{n}} & \dots & \frac{1}{\sqrt{n}} \\ \frac{1}{\sqrt{n-1}} & -\frac{1}{\sqrt{n-1}} & \dots & -\frac{1}{\sqrt{n-1}} \\ \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \vdots & c_{ij} \end{pmatrix} \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{pmatrix} \text{ 记 } Y = AX$$

$$\text{即 } Y_1 = \sqrt{n}\bar{X} \quad Y_2 = \frac{1}{\sqrt{n-1}}(X_1 - \bar{X}) \quad Y_i = \sum_{j=1}^n C_{ij} X_j \quad (i=3, 4, \dots, n)$$

$$\text{取 } A \text{ 为正交阵 正交变换保距 } \sum_{i=1}^n X_i^2 = \sum_{i=1}^n Y_i^2$$

$$(n-1)S^2 = \sum_{i=1}^n (X_i - \bar{X})^2 = \sum_{i=1}^n X_i^2 - n\bar{X}^2 = \sum_{i=2}^n Y_i^2$$

$$\text{由定理 2.2.3, } \frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$$

$$X_1 - \bar{X} = \frac{n-1}{n} X_1 - \sum_{i=2}^n \frac{1}{n} X_i \sim N(0, \frac{n-1}{n} \sigma^2) \left(\frac{(n-1)^2}{n^2} + \frac{n-1}{n^2} = \frac{n-1}{n} \right)$$

$$\Rightarrow Y_2 = \frac{1}{\sqrt{n-1}}(X_1 - \bar{X}) \sim N(0, \sigma^2)$$

$$Z = \frac{X_1 - \bar{X}}{S} = \frac{\frac{1}{\sqrt{n-1}} Y_2}{\frac{1}{\sqrt{n-1}} \sqrt{\sum_{i=2}^n Y_i^2}} = \frac{(n-1) Y_2}{\sqrt{n} \sqrt{\sum_{i=2}^n Y_i^2}}$$

$$\text{有 } \frac{1}{Z^2} = \frac{n}{(n-1)^2} \left(1 + \frac{\sum_{i=3}^n Y_i^2}{Y_2^2} \right) \quad \text{有 } \frac{\sum_{i=3}^n Y_i^2}{\sigma^2} \sim \chi_{n-2}^2$$

$$\text{取 } \eta = \frac{Y_2}{\sqrt{\frac{\sum_{i=3}^n Y_i^2}{\sigma^2(n-2)}}} = \frac{\sqrt{n-2} Y_2}{\sum_{i=3}^n Y_i^2} \sim t_{n-2}$$

$$\Rightarrow \frac{1}{Z^2} = \frac{n}{(n-1)^2} \left(1 + \frac{n-2}{\eta^2} \right), \quad Z \text{ 与 } \eta \text{ 同分布}$$

$$\text{即 } \eta = \sqrt{\frac{n-2}{\frac{(n-1)^2}{n^2} - 1}} \cdot \text{sgn}(Z)$$

$$\text{对 } x < 0, \quad F(x) = P(Z < x) = P\left(\eta < \frac{\sqrt{n-2}}{\sqrt{\frac{(n-1)^2}{n^2} - 1}}\right) = \int$$

由变量代换公式可得

$$f(z) = \frac{\Gamma(\frac{n-1}{2})}{\Gamma(\frac{n-2}{2}) \sqrt{\pi}} \left[\frac{(n-1)^2}{n} - z^2 \right]^{\frac{n-4}{2}} \left[\frac{n}{(n-1)^2} \right]^{\frac{n-3}{2}}$$

D.



$$14. \quad Z = \frac{\sum_{i=1}^n X_i^2}{\sigma_i^2} + Z^2 \sum_{i=1}^n \frac{1}{\sigma_i^2} \rightarrow Z \sum_{i=1}^n \frac{X_i}{\sigma_i^2}$$

$$= \sum_{i=1}^n \frac{X_i^2}{\sigma_i^2} - \sum_{i=1}^n \frac{1}{\sigma_i^2} Z^2$$

例取 $A = \begin{pmatrix} \frac{1}{\sigma_1} & \frac{1}{\sigma_2} & \dots & \frac{1}{\sigma_n} \\ \sqrt{\frac{n}{\sum_{i=1}^n \frac{1}{\sigma_i^2}}} & \sqrt{\frac{n}{\sum_{i=1}^n \frac{1}{\sigma_i^2}}} & \dots & \sqrt{\frac{n}{\sum_{i=1}^n \frac{1}{\sigma_i^2}}} \\ \vdots & \vdots & \ddots & \vdots \end{pmatrix}$

为正交阵 (Schmidt 正交化保证)。

$$Y_1 = \sqrt{\frac{n}{\sum_{i=1}^n \frac{1}{\sigma_i^2}}} Z \Rightarrow Z = \frac{\sum_{i=1}^n \frac{X_i^2}{\sigma_i^2} - Y_1^2}{\sum_{i=2}^n Y_i^2}$$

由正交变换保模长 $\Rightarrow$

且 $Y_1, Y_2, \dots, Y_n \sim N(0, 1)$ i.i.d.

$$\therefore Z \sim \chi_{n-1}^2.$$

□.

$$16. \quad Y = X_1 + \dots + X_n \sim \chi_{mn}^2$$

$$f(\bar{X}=x) = f(\sum X_i = nX) = f(Y=nX) = \begin{cases} \frac{n}{2^{\frac{mn}{2}} \Gamma(\frac{mn}{2})} (nX)^{\frac{mn}{2}-1} e^{-\frac{nX}{2}}, & x > 0 \\ 0, & x \leq 0 \end{cases}$$

□



20. 令 $\vec{Y} = \vec{X} - \vec{\mu}$

由 $\text{Cov}(\vec{X}) = \vec{\Sigma}_{n \times n} > 0$

$\therefore \exists A$ s.t. $\vec{\Sigma}_{n \times n} = A \cdot A^T$

令 $\vec{Z} = A^{-1} \vec{Y}$

由 $\vec{Y} \sim N(0, \vec{\Sigma}_{n \times n})$

$\Rightarrow \vec{Z} \sim N(0, A^{-1}(\vec{\Sigma}_{n \times n})(A^{-1})^T)$
 $\sim N(0, I_{n \times n})$

$\Rightarrow Z_i \sim N(0, 1)$ i.i.d. $i = 1, 2, \dots, n$

$$U = (\vec{X} - \vec{\mu})^T \vec{\Sigma}^{-1} (\vec{X} - \vec{\mu})$$

$$= \vec{Y}^T (A A^T)^{-1} \vec{Y}$$

$$= (A^{-1} \vec{Y})^T (A^{-1} \vec{Y})$$

$$= \vec{Z}^T \vec{Z}$$

$$= \sum_{i=1}^n Z_i^2 \Rightarrow U \sim \chi_n^2$$

□

28. 设 X_1 pdf $f_1(x_1)$, X_2 pdf $f_2(x_2)$ 且令 pdf $f(x_1, x_2)$, X_1, X_2 独立

$$f(x_1, x_2) = f_1(x_1) \cdot f_2(x_2) = \frac{\lambda^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} \cdot x_1^{\alpha_1 - 1} x_2^{\alpha_2 - 1} e^{-\lambda(x_1 + x_2)}$$

令 $\begin{cases} y_1 = x_1 + x_2 \\ y_2 = \frac{x_1}{x_1 + x_2} \end{cases} \Rightarrow \begin{cases} x_1 = y_1 y_2 \\ x_2 = y_1 - y_1 y_2 \end{cases} \quad \left| \frac{\partial(x_1, x_2)}{\partial(y_1, y_2)} \right| = y_1 \neq 0$

$$\hat{f}(y_1, y_2) = \frac{\lambda^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} \cdot (y_1 y_2)^{\alpha_1 - 1} (y_1 - y_1 y_2)^{\alpha_2 - 1} e^{-\lambda y_1} \cdot y_1$$

$$= \frac{\lambda^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1 + \alpha_2)} e^{-\lambda y_1} y_1^{\alpha_1 + \alpha_2 - 1} \frac{1}{B(\alpha_1, \alpha_2)} y_2^{\alpha_1 - 1} (1 - y_2)^{\alpha_2 - 1}$$

$$= \hat{f}_1(y_1) \hat{f}_2(y_2)$$

$\hat{f}(y_1, y_2)$ 可分离 $\Rightarrow Y_1$ 与 Y_2 相互独立

且 $Y_1 \sim \Gamma(\alpha_1 + \alpha_2, \lambda)$, $Y_2 \sim B(\alpha_1, \alpha_2)$.

□.



35. $X_1, \dots, X_n \stackrel{i.i.d.}{\sim} U(0, 1)$

$$E X_1 = \frac{1}{2} \quad D X_1 = \frac{1}{12}$$

$$\sqrt{n} \left(\frac{\bar{X} - \frac{1}{2}}{\sqrt{\frac{1}{12n}}} \right) \rightarrow N(0, 1)$$

$$\Rightarrow \bar{X} \text{ 的渐近分布为 } N\left(\frac{1}{2}, \frac{1}{12n}\right)$$

□

