

30. $H_0: \sigma^2 \leq \sigma_0^2 \longleftrightarrow H_1: \sigma^2 > \sigma_0^2$

记 $\theta = (\mu, \sigma^2)$ θ 的似然函数为

$$f(x, \theta) = (2\lambda\sigma^2)^{-\frac{n}{2}} \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2\right\}.$$

$$\Theta = \{\theta = (\mu, \sigma^2) : -\infty < \mu < +\infty, \sigma^2 > 0\}$$

$$\Theta_0 = \{\theta = (\mu, \sigma^2) : -\infty < \mu < +\infty, \sigma^2 \leq \sigma_0^2\}$$

在 Θ 上, μ, σ^2 的 MLE: $\hat{\mu} = \bar{x}, \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$

故有 $\sup_{\theta \in \Theta} f(x, \theta) = \left(\frac{2\lambda e}{n}\right)^{-\frac{n}{2}} \left(\sum_{i=1}^n (x_i - \bar{x})^2\right)^{-\frac{n}{2}}$

$$L_{\Theta_0}(x) = \sup_{\sigma^2 \leq \sigma_0^2} (2\lambda\sigma^2)^{-\frac{n}{2}} \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \bar{x})^2\right\}$$

$$\begin{aligned} \ln g(\sigma^2) &= \ln \left[(2\lambda\sigma^2)^{-\frac{n}{2}} \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \bar{x})^2\right\} \right] \\ &= -\frac{n}{2} (\ln(2\lambda) + \ln \sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \bar{x})^2 \end{aligned}$$

求得 $\Rightarrow$ ~~$L_{\Theta_0}(x) =$~~ $\hat{\mu}' = \bar{x}, \hat{\sigma}^2 = \begin{cases} \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2, & \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2 \leq \sigma_0^2 \\ \sigma_0^2, & \text{else} \end{cases}$

$$\therefore \lambda(x) = \begin{cases} 1 & \sigma_0^2 > \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n} \\ \left(\frac{e}{n}\right)^{-\frac{n}{2}} \left[\frac{1}{\sigma_0^2} \sum_{i=1}^n (x_i - \bar{x})^2\right]^{-\frac{n}{2}} \exp\left\{\frac{1}{\sigma_0^2} \sum_{i=1}^n (x_i - \bar{x})^2\right\}, & \sigma_0^2 < \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n} \end{cases}$$

$$\text{令 } \xi = \frac{1}{\sigma_0^2} \sum_{i=1}^n (x_i - \bar{x})^2, \quad g(\xi) = \xi^{-\frac{n}{2}} e^{\frac{\xi}{2}} \left(\frac{e}{n}\right)^{-\frac{n}{2}}$$

$\xi > 0$ 时 $g(\xi)$ 在 $(0, n) \downarrow$, $(n, +\infty) \nearrow$

则 $\lambda(x)$ 在 $(0, n)$ 中 $\downarrow$, 在 $(n, +\infty)$ 中 $\nearrow$

故 $p(g(\xi) \leq c) = p(\xi \leq k)$ 再 ξ 在 n 附近时服从 χ_{n-1}^2

故 $k = \chi_{n-1}^2(\alpha) \quad \therefore \quad \varphi(x) = \begin{cases} 1, & \frac{\sum (x_i - \bar{x})^2}{\sigma_0^2} > \chi_{n-1}^2(\alpha) \\ 0, & \text{else.} \end{cases}$

□



$$31. f(x, y, \theta) = (2\pi\sigma^2)^{-\frac{m+n}{2}} \exp\left\{-\frac{1}{2\sigma^2} \left[\sum_{i=1}^m (x_i - \mu_1)^2 + \sum_{i=1}^n (y_i - \mu_2)^2 \right]\right\}$$

$$(H) = \{\theta = (\mu_1, \mu_2, \sigma^2) : -\infty < \mu_1, \mu_2 < +\infty, \sigma^2 > 0\}$$

由指数族可求得

$$(H)_1 \text{ 上, } \hat{\mu}_1 = \bar{x}, \hat{\mu}_2 = \bar{y}, \hat{\sigma}^2 = \frac{\sum_{i=1}^m (x_i - \bar{x})^2 + \sum_{j=1}^n (y_j - \bar{y})^2}{m+n}$$

$$(H)_0 \text{ 上, } \hat{\mu}' = \frac{m\bar{x} + n\bar{y}}{m+n}, \hat{\sigma}'^2 = \frac{\sum_{i=1}^m (x_i - \hat{\mu}')^2 + \sum_{j=1}^n (y_j - \hat{\mu}')^2}{m+n}$$

代入, $\Rightarrow$

$$\lambda(\bar{x}) = \frac{\cancel{L_0(x)} L_1(x)}{L_0(x)} = \left[\frac{\sum_{i=1}^m (x_i - \bar{x})^2 + \sum_{j=1}^n (y_j - \bar{y})^2}{\sum_{i=1}^m (x_i - \hat{\mu}')^2 + \sum_{j=1}^n (y_j - \hat{\mu}')^2} \right]^{-\frac{m+n}{2}}$$

$$\sum_{i=1}^m (x_i - \hat{\mu}')^2 = \sum_{i=1}^m (x_i - \bar{x})^2 + m(\bar{x} - \hat{\mu}')^2 = \sum_{i=1}^m (x_i - \bar{x})^2 + \frac{mn(\bar{x} - \bar{y})^2}{(m+n)^2}$$

$$\text{同理 } \sum_{j=1}^n (y_j - \hat{\mu}')^2 = \sum_{j=1}^n (y_j - \bar{y})^2 + \frac{m^2n(\bar{x} - \bar{y})^2}{(m+n)^2}$$

$$\Rightarrow \lambda(x, y) = \left(1 + \frac{mn}{m+n} \cdot \frac{(\bar{x} - \bar{y})^2}{\sum_{i=1}^m (x_i - \bar{x})^2 + \sum_{j=1}^n (y_j - \bar{y})^2} \right)^{-\frac{m+n}{2}}$$

$$T = \frac{\bar{x} - \bar{y}}{\sqrt{\frac{1}{m} + \frac{1}{n}} \sqrt{\sum_{i=1}^m (x_i - \bar{x})^2 + \sum_{j=1}^n (y_j - \bar{y})^2 / (m+n-2)}} \Rightarrow T \sim t_{m+n-2} \quad (\text{上下独立}).$$

$$P(|T| > d) = \alpha \Rightarrow d = t_{m+n-2}(\frac{\alpha}{2})$$

$$\therefore f(x) = \begin{cases} 0, & |T| \leq t_{m+n-2}(\frac{\alpha}{2}) \\ 1, & |T| > t_{m+n-2}(\frac{\alpha}{2}) \end{cases}$$

□.



34. $f(x, y, \theta) \rightarrow$

$$\sum_{i=1}^m (x_i - \bar{x})^2 = A, \quad \sum_{i=1}^n (y_i - \bar{y})^2 = B.$$

$$f(x, y, \theta) = (2\pi)^{-\frac{m+n}{2}} (\sigma_1^2)^{-\frac{m}{2}} (\sigma_2^2)^{-\frac{n}{2}} \exp \left\{ -\frac{1}{2\sigma_1^2} \sum_{i=1}^m (x_i - \mu_1)^2 - \frac{1}{2\sigma_2^2} \sum_{i=1}^n (y_i - \mu_2)^2 \right\}$$

由 $\hat{\theta}_M$:
 (A) $\hat{\mu}_1 = \bar{x}, \hat{\mu}_2 = \bar{y}, \hat{\sigma}_1^2 = \frac{A}{m}, \hat{\sigma}_2^2 = \frac{B}{n}$

(A)₀ $\hat{\mu}_1 = \bar{x}, \hat{\mu}_2 = \bar{y}$ 设 $\sigma_1^2 = c\sigma_2^2 = c\sigma^2$

$$\hat{\sigma}^2 = \frac{\sum (x_i - \bar{x})^2 + \sum (y_i - \bar{y})^2}{c(m+n)}$$

$$L_{(A)_0}(x) = (2\pi e)^{-\frac{m+n}{2}} (\hat{\sigma}^2)^{-\frac{m+n}{2}} c^{-\frac{m}{2}}$$

$$\lambda(x) = \frac{(\hat{\sigma}^2)^{-\frac{m}{2}} (\hat{\sigma}^2)^{-\frac{n}{2}}}{(\hat{\sigma}^2)^{-\frac{m+n}{2}}} c^{-\frac{m}{2}}$$

$$\lambda(x) = \left(\frac{m+n}{n}\right)^{-\frac{n}{2}} \left(\frac{m+n}{m}\right)^{-\frac{n}{2}} c^{-\frac{n}{2}} \left(\frac{A}{A+CB}\right)^{-\frac{m}{2}} \left(\frac{B}{A+CB}\right)^{-\frac{n}{2}}$$

$$\hat{\lambda} F = \frac{A/\sqrt{c}m}{B/n} \quad \text{即 } \theta \in (A)_0 \quad F \sim F_{m, n-1}$$

$$\Rightarrow \lambda(x, y) = c(m, n) \left(1 + \frac{n\sqrt{c}}{mF}\right) \left(c + \frac{m\sqrt{c}}{n} F\right), \quad \text{下证 } F \text{ 与 } \lambda \text{ 成 } 1 \text{ 对 } \lambda$$

由 $\lambda \leq c \Leftrightarrow K_1 \leq F \leq K_2$ 取 $\begin{cases} P(F > K_2) = \frac{\alpha}{2} \\ P(F < K_1) = \frac{\alpha}{2} \end{cases}$

$$g(x) = \begin{cases} 0 & F_{m-1, n-1}(1-\frac{\alpha}{2}) \leq F \leq F_{m-1, n-1}(\frac{\alpha}{2}) \\ 1 & \text{else} \end{cases}$$

□.



28. 例 5.3.1

$$L(x, p) = \binom{n}{\sum x_i} p^{\sum x_i} (1-p)^{n-\sum x_i} \quad x_i \in \{0, 1\}, i=1, \dots, n.$$

$$\Theta = \{p: p \in [0, 1]\} \quad \Theta_0 = \{p_0\}$$

$$\Rightarrow L_{\Theta}(x, p) = \binom{n}{\sum x_i} (\bar{x})^{\sum x_i} (1-\bar{x})^{n-\sum x_i}$$

$$L_{\Theta_0}(x, p) = \binom{n}{\sum x_i} p_0^{\sum x_i} (1-p_0)^{n-\sum x_i}$$

$$\therefore \lambda(x) = \left(\frac{\bar{x}}{p_0}\right)^{\sum x_i} \left(\frac{1-\bar{x}}{1-p_0}\right)^{n-\sum x_i}$$

Θ 自然参数 λ Θ_0 自然参数 λ_0 .

$$\therefore \lambda(x) \propto \sum x_i^2$$

$$\Rightarrow \varphi(x) = \begin{cases} 0, & \lambda(x) \leq e^{\frac{\chi_1^2(\alpha)}{2}} \\ 1, & \lambda(x) > \chi_1^2(\alpha) e^{\frac{\chi_1^2(\alpha)}{2}} \end{cases}$$

12



41. 由因子分解定理 $f(x, \theta) = g(x(x), \theta) h(x)$

$$\Rightarrow f(x, \theta_1) = g(x(x), \theta_1) h(x)$$

$$f(x, \theta_0) = g(x(x), \theta_0) h(x)$$

$$\Rightarrow \frac{f(x, \theta_1)}{f(x, \theta_0)} = \frac{g(x(x), \theta_1)}{g(x(x), \theta_0)} \triangleq h(x(x))$$

故由 NP 定理知是

$$g(x) = \int_0^1 r; \begin{matrix} h(x(x)) > c \\ h(x(x)) = c \\ h(x(x)) < c \end{matrix} \quad \text{为 } T \text{ 的函数,}$$

□.

42. (1) $X_1, \dots, X_n$ 联合 p.d.f $f(\vec{x}, \theta) = \frac{1}{\theta^n} \mathbb{I}[X_{(n)} < \theta]$

$$\Rightarrow \lambda(x) = \frac{f(x, \theta_1)}{f(x, \theta_0)} = \begin{cases} \frac{1}{2^n} & 0 < X_{(n)} < 1 \\ +\infty & X_{(n)} \geq 1 \end{cases} \quad (c \triangleq \frac{\theta_1}{\theta_0} = 1/2)$$

$$\Rightarrow g(x) = \int_0^1; \begin{matrix} \lambda(x) > c \\ \lambda(x) \leq c \end{matrix} = \int_0^1; \begin{matrix} X_{(n)} > c \\ X_{(n)} \leq c \end{matrix}$$

$$X_{(n)} \text{ p.d.f } g_{\theta}(x) = \frac{n x^{n-1}}{\theta^n} \mathbb{I}[0 < x < \theta]$$

$$\Rightarrow E_{\theta_0}[g(x)] = \int_0^{\theta_0} \frac{n x^{n-1}}{\theta_0^n} dx = 1 - \frac{c^n}{\theta_0^n} = 1 - c^n = \alpha$$

$$\Rightarrow c = \sqrt[n]{1-\alpha}$$

$$\therefore g(x) = \begin{cases} 1 & X_{(n)} > \sqrt[n]{1-\alpha} \\ 0 & X_{(n)} \leq \sqrt[n]{1-\alpha} \end{cases} \quad \text{为拒绝域 UMP}$$

b



$$(2). \quad \lambda(x) = \begin{cases} z^n, & 0 < x_{(n)} \leq \frac{1}{z} \\ 0, & \text{else} \end{cases} \quad (z > \frac{0}{0} = 0)$$

由 $\lambda(x) \downarrow$

$$\therefore \varphi(x) = \begin{cases} 1 & x_{(n)} < c \\ 0 & x_{(n)} \geq c \end{cases}$$

$$E[\varphi(x)] = \int_0^c \frac{n x^{n-1}}{\theta_0^n} dx = \frac{c^n}{\theta_0^n} = c^n = \alpha$$

$$\Rightarrow c = \eta \sqrt[n]{\alpha}$$

$$\Rightarrow \varphi(x) = \begin{cases} 1 & x_{(n)} < \eta \sqrt[n]{\alpha} \\ 0 & x_{(n)} \geq \eta \sqrt[n]{\alpha} \end{cases}$$

→ 检验 U_{MPT} .

b

$$43. \quad f(x, \mu) = (2\pi)^{-\frac{n}{2}} e^{-\frac{n}{2}\mu^2} \exp\{\mu \cdot \sum x_i\} \cdot \exp\{-\frac{1}{2} \sum x_i^2\}$$

$$c(\mu) = e^{-\frac{n}{2}\mu^2} \quad Q(\mu) = \mu \quad h(x) = \exp\{-\frac{1}{2} \sum x_i^2\}, \quad T(x) = \sum x_i$$

$Q(\mu)$ 关于 μ 单调 $\Rightarrow$

$$U_{MPT} \text{ 形式 } \varphi(x) = \begin{cases} 1 & T(x) < nc \\ 0 & T(x) \geq nc \end{cases}$$

$$E_{\mu_0}(\varphi(x)) = P_{\mu_0}(T(x) < nc) = P_{\mu_0}(\bar{x} < c) = \alpha$$

$$P_{\mu_0}(\sqrt{n}(\bar{x} - \mu_0) < \sqrt{n}(c - \mu_0)) = \alpha \Rightarrow$$

$$\sqrt{n}(\bar{x} - \mu_0) = -u_\alpha$$

$$\Rightarrow c = \mu_0 - \frac{u_\alpha}{\sqrt{n}}$$

$$\therefore \varphi(x) = \begin{cases} 1 & , \bar{x} < \mu_0 - \frac{u_\alpha}{\sqrt{n}} \\ 0 & , \text{else} \end{cases}$$

□.



46. $H_0: p = \frac{1}{2} \leftrightarrow H_1: p < \frac{1}{2}$

$$f(x_1, \dots, x_n, p) = \binom{n}{\sum x_i} p^{\sum x_i} (1-p)^{n-\sum x_i}$$

$$= (1-p)^n \exp \left\{ \ln \left(\frac{p}{1-p} \right) \cdot \sum x_i \right\}$$

为充分统计量 $T(x) = \sum x_i$ $Q(p) = \ln \frac{p}{1-p}$ 关于 p 严格递增

$$\therefore g(x) = \begin{cases} 1 & , \sum x_i < c \\ r & , \sum x_i = c \\ 0 & , \sum x_i > c \end{cases}$$

由 $\sum x_i \sim b(n, p)$

$$\begin{aligned} E_{p_0}[g(x)] &= \sum_{i=0}^{c-1} \binom{n}{i} p_0^i (1-p_0)^{n-i} + r \binom{n}{c} p_0^c (1-p_0)^{n-c} \\ &= \frac{1}{2^n} \left[\sum_{i=0}^{c-1} \binom{n}{i} + r \binom{n}{c} \right] \end{aligned}$$

由 $\frac{1}{2^n} \sum_{i=1}^{c-1} \binom{n}{i} \leq \alpha < \frac{1}{2^n} \sum_{i=1}^c \binom{n}{i}$ 确定.

$$r = \frac{\alpha - \frac{1}{2^n} \sum_{i=0}^{c-1} \binom{n}{i}}{\binom{n}{c} \frac{1}{2^n}}$$

□.

